

Models In Epidemiology And Biostatistics

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Example 2 with Annotation

```
. use Session_4_Example_2.dta
```

```
. cs dis expo,by(age gender) or
```

age gender	OR	[95% Conf. Interval]		M-H Weight	
0 0	2.181717	1.256808	3.785856	10.35	(Cornfield)
0 1	1.962957	1.327944	2.901236	18.43281	(Cornfield)
1 0	3.305753	1.953334	5.592873	10.37538	(Cornfield)
1 1	3.379872	1.92368	5.934744	7.374564	(Cornfield)
Crude	.4575415	.3940612	.5312483		
M-H combined	2.535571	1.981625	3.244369		
Test of homogeneity (M-H) chi2(3) = 3.851 Pr>chi2 = 0.2780					
Test that combined OR = 1:					
Mantel-Haenszel				chi2(1) =	57.36
				Pr>chi2 =	0.0000

For the young, male and female OR estimates are close and about 2. For the old, male and female OR estimates are close and about 3.3. Modification by age ? Crude and adjusted are very different so there is some kind of confounding.

```
. cs dis expo,by(gender) or
```

gender	OR	[95% Conf. Interval]		M-H Weight	
0	2.742677	1.871612	4.01882	20.74615	(Cornfield)
1	2.34375	1.707093	3.217633	26.35914	(Cornfield)
Crude	.4575415	.3940612	.5312483		
M-H combined	2.519446	1.971614	3.219498		
Test of homogeneity (M-H) chi2(1) = 0.386 Pr>chi2 = 0.5343					
Test that combined OR = 1:					
Mantel-Haenszel				chi2(1) =	57.06
				Pr>chi2 =	0.0000

This 'one-at-a-time' analysis does not display the age difference seen in the joint analysis but does reveal a confounding issue.

```
. cs dis expo,by(age) or
```

age	OR	[95% Conf. Interval]		M-H Weight	
0	.4168153	.3369634	.5155913	129.4351	(Cornfield)
1	.5023695	.4072843	.6196537	121.6115	(Cornfield)
Crude	.4575415	.3940612	.5312483		
M-H combined	.4582593	.3946601	.5321074		
Test of homogeneity (M-H) chi2(1) = 1.498 Pr>chi2 = 0.2210					
Test that combined OR = 1:					
Mantel-Haenszel				chi2(1) =	107.15
				Pr>chi2 =	0.0000

This 'one-at-a-time' analysis misses the confounding and misses the modification.

```
. cs dis expo if age==0,by(gender) or
```

gender	OR	[95% Conf. Interval]		M-H Weight	
0	2.181717	1.256808	3.785856	10.35	(Cornfield)
1	1.962957	1.327944	2.901236	18.43281	(Cornfield)
Crude	.4168153	.3369634	.5155913		
M-H combined	2.041621	1.481015	2.814431		
Test of homogeneity (M-H) chi2(1) = 0.093 Pr>chi2 = 0.7605					
Test that combined OR = 1:					
Mantel-Haenszel				chi2(1) =	19.42
				Pr>chi2 =	0.0000

For the young, adjusting for gender looks promising.

```
. cs dis expo if age==1,by(gender) or
```

gender	OR	[95% Conf. Interval]		M-H Weight	
0	3.305753	1.953334	5.592873	10.37538	(Cornfield)
1	3.379872	1.92368	5.934744	7.374564	(Cornfield)
Crude	.5023695	.4072843	.6196537		
M-H combined	3.336548	2.257874	4.930545		
Test of homogeneity (M-H) chi2(1) = 0.003 Pr>chi2 = 0.9553					
Test that combined OR = 1:					
Mantel-Haenszel				chi2(1) =	40.27
				Pr>chi2 =	0.0000

For the old, adjusting for gender also seems reasonable.

There appears to be age modification but there is no test available here. The confidence intervals for the assumed common ORs overlap.

Now we can try a model based approach

```
. logit dis expo age gender ga ge ae gae
```

dis	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
expo	.7801121	.2846032	2.74	0.006	.2223	1.337924
age	.0680535	.3690482	0.18	0.854	-.6552677	.7913746
gender	2.952439	.2795989	10.56	0.000	2.404435	3.500443
ga	-.0840535	.3901241	-0.22	0.829	-.8486828	.6805757
ge	-.1056603	.3478914	-0.30	0.761	-.7875148	.5761943
ae	.4155523	.3931613	1.06	0.291	-.3550297	1.186134
gae	.1278339	.5281787	0.24	0.809	-.9073773	1.163045
_cons	-2.944439	.2649065	-11.12	0.000	-3.463646	-2.425232

There is no evidence that age modification is modified by gender [p=0.809].

From the classic analysis, for the women, we got $3.31 / 2.18 = 1.5$ and for the men, we get $3.38 / 1.96 = 1.7$.

Similar numbers can be computed for the above model. $\exp(0.4155) = 1.5$ $\exp(0.4155+0.127) = 1.7$

```
. logit dis expo age gender ga ge ae
```

	dis	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
	expo	.7433503	.2377981	3.13	0.002	.2772746	1.209426
	age	.0055253	.2633948	0.02	0.983	-.5107191	.5217697
	gender	2.916984	.2351269	12.41	0.000	2.456144	3.377824
	ga	-.0141751	.2622424	-0.05	0.957	-.5281608	.4998107
	ge	-.0504734	.260819	-0.19	0.847	-.5616693	.4607224
	ae	.4865294	.2617646	1.86	0.063	-.0265198	.9995786
	_cons	-2.912659	.2268167	-12.84	0.000	-3.357212	-2.468107

0.4865 is the estimate of the log of the ratio of assumed common ratio of ORs [old over young : adjusted for gender] . This is interesting and is supported by the classic analysis.

-0.0504 is this estimate [male over female adjusted for age]. It is near zero.

```
. logit dis expo age gender ga ae
```

	dis	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
	expo	.7100288	.1628327	4.36	0.000	.3908825	1.029175
	age	-.0017126	.2597452	-0.01	0.995	-.5108038	.5073787
	gender	2.884892	.1654006	17.44	0.000	2.560713	3.209071
	ga	-.0060814	.2583501	-0.02	0.981	-.5124383	.5002755
	ae	.4959979	.2566618	1.93	0.053	-.0070499	.9990458
	_cons	-2.883975	.1701056	-16.95	0.000	-3.217376	-2.550574

We will take the step to delete the ga term. It is instructive to interpret it though.

```
. logit dis expo age gender ae
```

Logistic regression	Number of obs	=	3,814
	LR chi2(4)	=	840.37
	Prob > chi2	=	0.0000
Log likelihood = -1710.7169	Pseudo R2	=	0.1972

	dis	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
	expo	.7084344	.1480128	4.79	0.000	.4183347	.9985341
	age	-.0071415	.119512	-0.06	0.952	-.2413807	.2270977
	gender	2.882402	.1270758	22.68	0.000	2.633338	3.131466
	ae	.5004742	.1724266	2.90	0.004	.1625242	.8384242
	_cons	-2.881753	.1414169	-20.38	0.000	-3.158925	-2.604581

Now we have the model that provides for age modification adjusted for gender.

```
. logit dis expo age gender ae, or
```

	dis	Odds Ratio	Std. Err.	z	P> z	[95% Conf. Interval]	
	expo	2.030809	.3005857	4.79	0.000	1.519429	2.7143
	age	.9928839	.1186615	-0.06	0.952	.7855425	1.254952
	gender	17.85711	2.269207	22.68	0.000	13.92015	22.90753
	ae	1.649503	.2844183	2.90	0.004	1.176477	2.31272
	_cons	.0560365	.0079245	-20.38	0.000	.0424714	.0739341

```
. disp 2.030809*1.649503
3.3498255
```

Taking the exponent of the coefficient estimates gives us numbers very close to the classic analysis. Perhaps most crucially this model provides clear evidence [p=0.004] that age modifies after adjusted for gender.

```
. logit dis expo age ae
```

dis	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
expo	-.8751117	.1085975	-8.06	0.000	-1.087959	-.6622645
age	-.0056374	.1061829	-0.05	0.958	-.213752	.2024772
ae	.1866923	.1525511	1.22	0.221	-.1123025	.485687
_cons	-.6968995	.075047	-9.29	0.000	-.843989	-.54981

```
. logit dis expo age ae,or
```

dis	Odds Ratio	Std. Err.	z	P> z	[95% Conf. Interval]	
expo	.4168155	.0452651	-8.06	0.000	.3369034	.5156822
age	.9943785	.105586	-0.05	0.958	.8075486	1.224432
ae	1.205256	.1838632	1.22	0.221	.8937739	1.625291
_cons	.4981273	.037383	-9.29	0.000	.4299919	.5770594

We can now compare 1.6495 [assessing age modification adjusted for gender] with 1.2052 [assessing age modification without adjustment for gender] $1.6495/1.2052 = 1.37$. Relative to the 'crude', the adjusted is 37% more than the 'crude' here. We have evidence that age modification is confounded by gender. This confounding is identified by comparing the two estimates and is identified by noting there is statistical significance with the adjusted but there is no statistical significance with the 'crude'.